

Problem 9.1:

(a)

$$\begin{aligned}
h(X) &= - \int_0^{\infty} \lambda e^{-\lambda x} [\ln \lambda - \lambda x] dx \text{ (in nats)} \\
&= - \ln \lambda \underbrace{\left[\int_0^{\infty} \lambda e^{-\lambda x} dx \right]}_{=1} + \lambda \underbrace{\left[\int_0^{\infty} x \lambda e^{-\lambda x} dx \right]}_{=E[X]=1/\lambda} \\
&= 1 - \ln \lambda \text{ nats} \\
&= (1 - \ln \lambda) / \ln 2 \text{ bits.}
\end{aligned}$$

(b)

$$\begin{aligned}
h(X) &= - \int_{-\infty}^{\infty} \frac{1}{2} \lambda e^{-\lambda|x|} \left[\ln \frac{\lambda}{2} - \lambda|x| \right] dx \\
&= - \ln(\lambda/2) + \lambda E[|X|] \\
&= - \ln(\lambda/2) + 2\lambda \int_0^{\infty} (x/2) \lambda e^{-\lambda x} dx \\
&= - \ln(\lambda/2) + (2\lambda) \frac{1}{2\lambda} \\
&= 1 - \ln(\lambda/2) \text{ nats.}
\end{aligned}$$

(c) $Y = X_1 + X_2$, where $X_1 \sim \mathcal{N}(\mu_1, \sigma_1^2)$ and $X_2 \sim \mathcal{N}(\mu_2, \sigma_2^2)$.

$$\implies Y \sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

$$\implies h(Y) = \frac{1}{2} \log_2(2\pi e(\sigma_1^2 + \sigma_2^2))$$

Problem 9.3:For $\rho = 0$: X, Y are uncorrelated Gaussian $\implies X, Y$ are independent $\implies I(X; Y) = 0$.For general case ($-1 \leq \rho \leq 1$):

$$I(X; Y) = h(X) + h(Y) - h(X, Y)$$

where

$$h(X) = h(Y) = \frac{1}{2} \log_2(2\pi e\sigma^2)$$

$$\begin{aligned}
h(X, Y) &= \frac{1}{2} \log_2 \left[(2\pi e)^2 \begin{vmatrix} \sigma^2 & \rho\sigma^2 \\ \rho\sigma^2 & \sigma^2 \end{vmatrix} \right] \text{ from (9.94)} \\
&= \frac{1}{2} \log_2 \left((2\pi e)^2 (1 - \rho^2) \sigma^4 \right)
\end{aligned}$$

$$\implies I(X; Y) = \frac{1}{2} \log_2 \left(\frac{1}{1 - \rho^2} \right)$$

For $\rho = \pm 1$: In this case, $I(X; Y) = +\infty$. This is because $X = \pm Y$ with probability one.

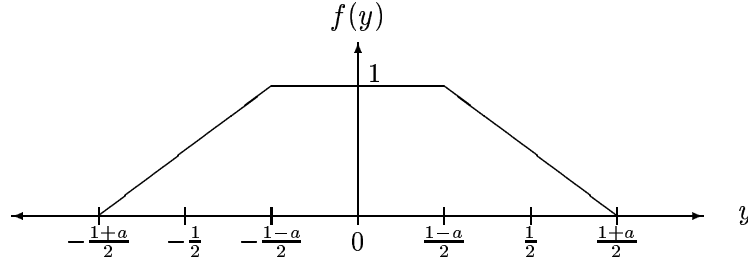
It is important to note that $I(X;Y)$ is independent of σ^2 . Also note that there is a one-to-one correspondence between $|\rho|$ and $I(X;Y)$. The higher the (magnitude of) correlation, the higher the mutual information. Zero correlation implies zero mutual information. The correlation and the mutual information are both measures of dependency.

Problem 9.4:

X is uniformly distributed on $[-1/2, +1/2]$. Z is uniformly distributed on $[-a/2, +a/2]$. $Y = X + Z$

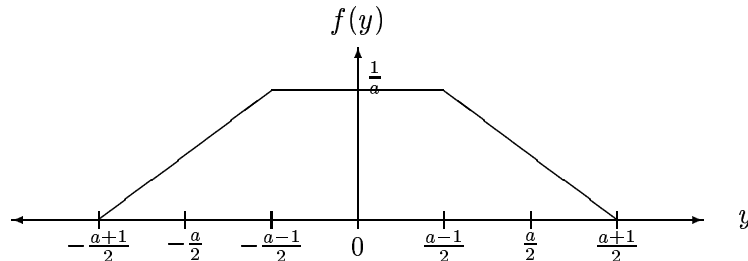
(a)

Case 1: $a \leq 1$: Here, the pdf of Y looks like:



$$\begin{aligned}
 h(Y) &= -2 \int_0^{(1-a)/2} 1 \log_2 1 dy - 2 \int_{(1-a)/2}^{(1+a)/2} -\frac{1}{a} \left(y - \frac{1+a}{2} \right) \log_2 \left[-\frac{1}{a} \left(y - \frac{1+a}{2} \right) \right] dy \\
 &= 0 + \frac{2}{a} \int_{-a}^0 x \log_2 \left[-\frac{x}{a} \right] dx && \left\{ x = y - \frac{1+a}{2} \right\} \\
 &= -2a \int_0^1 t \log_2 t dt && \left\{ t = -\frac{x}{a}; dx = -adt \right\} \\
 &= \frac{-2a}{\ln 2} \left[\frac{t^2}{2} \ln t - \frac{t^2}{4} \right]_0^1 \\
 &= \frac{a}{2 \ln 2} \text{ bits} = \frac{a}{2} \text{ nats.}
 \end{aligned}$$

Case 2: $a > 1$: Here, the pdf of Y looks like:



$$\begin{aligned}
 h(Y) &= -2 \int_0^{(a-1)/2} \frac{1}{a} \log_2 \frac{1}{a} dy - 2 \int_{(a-1)/2}^{(a+1)/2} -\frac{1}{a} \left(y - \frac{a+1}{2} \right) \log_2 \left[-\frac{1}{a} \left(y - \frac{a+1}{2} \right) \right] dy \\
 &= \ln a \left(\frac{a-1}{a} \right) - 2 \int_0^1 \frac{x}{a} \log_2 \left[\frac{x}{a} \right] dx && \left\{ x = -y + \frac{a+1}{2} \right\}
 \end{aligned}$$

$$\begin{aligned}
&= \left(1 - \frac{1}{a}\right) \ln a - 2a \int_0^{\frac{1}{a}} t \log_2 t dt && \left\{ t = \frac{x}{a}; dx = a dt \right\} \\
&= \left(1 - \frac{1}{a}\right) \ln a - 2a \left[\frac{t^2}{2} \ln t - \frac{t^2}{4} \right]_0^{\frac{1}{a}} \\
&= \ln a + \frac{1}{2a} \text{ nats.}
\end{aligned}$$

Now, $h(Z) = \ln a$ nats. Thus

$$I(X; Y) = h(Y) - h(Z) = \begin{cases} \frac{a}{2} - \ln a \text{ nats} & \text{if } 0 \leq a \leq 1 \\ \frac{a}{2} \text{ nats} & \text{if } a > 1 \end{cases}$$

(b) If $a = 1$, then $X \in [-1/2, +1/2]$ and $Z \in [-1/2, +1/2]$. This implies that $Y \in [-1, +1]$. Since, $I(X; Y) = h(Y) - h(Z) = h(Y) - \log_2(1) = h(Y)$,

$$\max_{|X| \leq 1/2} I(X; Y) = \max_{|X| \leq 1/2} h(Y).$$

The maximum value of $h(Y)$ is achieved when Y is uniformly distributed on $[-1, +1]$. This occurs if X is discrete and $\Pr\{X = -1/2\} = \Pr\{X = +1/2\} = \frac{1}{2}$. Thus, the channel capacity is

$$C = \log_2 2 - \log_2 1 = 1 \text{ bit.}$$

For this special channel, we can find a simple one-dimensional code with $R = C$ and λ_{max} =maximum probability of error=0. To do this, we send one of the two codewords: $x_1 = -1/2$ or $x_2 = +1/2$. (Note that these satisfied the *amplitude* constraint. When we receive y , we decode x_1 if $y < 0$ and we decode x_2 if $y \geq 0$. This results in zero probability of error.

Problem 9.5:

The decay time is typically modeled as an exponential distribution, $f(x) = \lambda e^{-\lambda x}$, $x \geq 0$. From problem 9.1(a), we know $h(X) = \log_2(e/\lambda)$ bits.

Since the median is 80,

$$\int_0^{80} \lambda e^{-\lambda x} dx = \frac{1}{2} \Leftrightarrow (1 - e^{-80\lambda}) = \frac{1}{2}.$$

So $\lambda = (\ln 2)/80 = 0.00866$.

Three digits accuracy are equal to $\lceil \log_2 1000 \rceil = \lceil 9.97 \rceil = 10$ fractional bits.

Thus, to represent the r.v. to 10 fractional bits accuracy we need $\approx h(X) + 10 = \log_2 \frac{e}{\lambda} + 10 = 18.29$ bits.

Problem B:

The covariance matrix is:

$$K = \begin{bmatrix} 1 & \rho & 0 \\ \rho & 1 & \rho \\ 0 & \rho & 1 \end{bmatrix}$$

and $\det K = 1 - 2\rho^2$. Thus,

$$\begin{aligned}
h(X, Y, Z) &= \frac{1}{2} \log(2\pi e)^3 (\det K) \\
&= \frac{1}{2} \log(2\pi e)^3 (1 - 2\rho^2) \\
&= \frac{3}{2} \log(2\pi e) + \frac{1}{2} \log(1 - 2\rho^2).
\end{aligned}$$

Since $(\det K)$ must be positive, we need that $1 - 2\rho^2 > 0$ or $|\rho| < \sqrt{2}/2$.

Please let me know if you see a mistake in this solution.